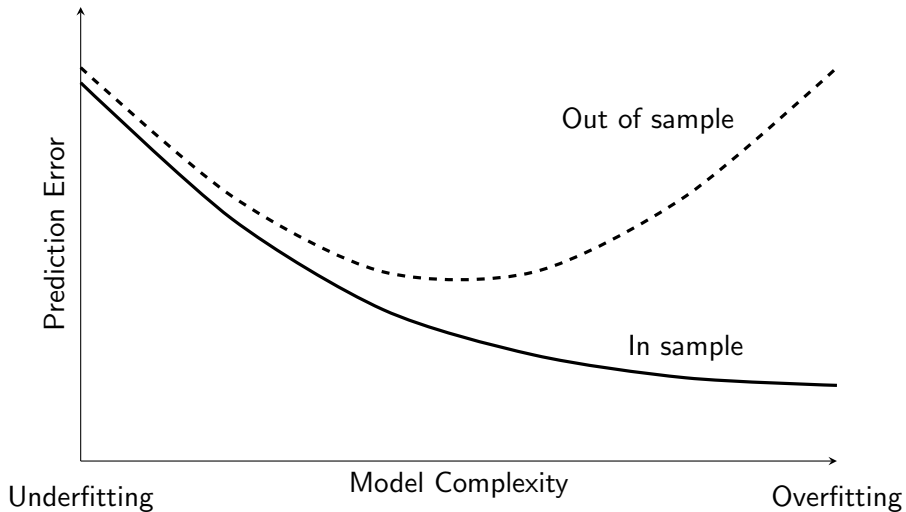


# The risk function of regularized ERM estimators tuned using CV

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## The basic tradeoff of supervised learning



# Introduction

- Standard supervised learning:
  - How to control model complexity?  
A: penalization / regularization / shrinkage.
  - How to find the optimal amount of regularization?  
A: Minimize cross-validation estimate of out-of-sample loss.
  - How to evaluate estimators?  
A: Expected out of sample loss.
    - Standard theory: Worst-case regret.
    - This talk: Risk function - more fine-grained picture!  
Taking into account data-dependent tuning!
- Key idea in our paper: We can approximate risk by
  - the mean squared error (MSE)
  - of shrinkage estimators in the normal means model
  - tuned using Stein's unbiased risk estimate (SURE).

## Review: Penalized ERM-estimation with tuning

- Penalized empirical risk minimization (ERM) estimator (in local coordinates):

$$\hat{\theta}_n^\lambda = \operatorname{argmin}_{\theta} \left[ \sum_{i=1}^n l(\theta/\sqrt{n}, Z_n^i) + \lambda \cdot \pi(\theta) \right].$$

- Cross-validated tuning parameter (definition uses **leave-one-out** estimator):

$$\lambda_n^* = \operatorname{argmin}_{\lambda} \sum_i l(\hat{\theta}_n^{\lambda, -i}/\sqrt{n}, Z_n^i).$$

- Risk function of the **tuned** estimator:

$$E \left[ l(\hat{\theta}_n^{\lambda_n^*}/\sqrt{n}, Z) \right],$$

for  $Z$  an independent draw.

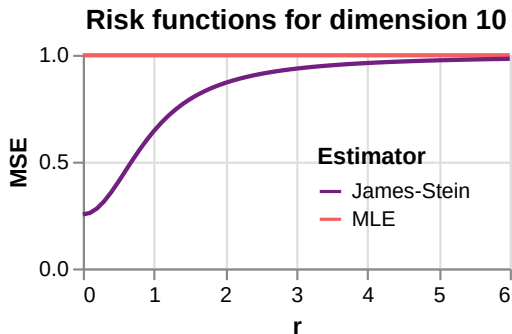
## Review: James-Stein shrinkage

- Shrinkage estimator: For  $\hat{\theta} \sim N(\theta, I_k)$ ,

$$\hat{\theta}^* = \left(1 - \frac{(k-2)}{\|X\|^2}\right) \cdot \hat{\theta}.$$

- Risk function (MSE):  $1 - \frac{1}{k} E \left[ \frac{(k-2)^2}{\|\hat{\theta}\|^2} \right]$ .

Denoting  $r = \|\theta_0\|$ :



## Review: Stein's unbiased risk estimate

- Suppose  $\hat{\theta} \sim N(\theta_0, \Sigma)$ . Let

$$\hat{\theta}^\lambda = \hat{\theta} + g^\lambda(\hat{\theta}),$$
$$g^\lambda(\theta) = \operatorname{argmin}_g \frac{1}{2} \|g\|^2 + \lambda \cdot \pi(\theta + g).$$

- Then

$$SURE(\lambda, \hat{\theta}, \Sigma) = \operatorname{trace}(\Sigma) + \underbrace{\|g^\lambda(\hat{\theta})\|^2}_{\text{In-sample loss}} + 2 \cdot \underbrace{\operatorname{trace}(\nabla g^\lambda(\hat{\theta}) \cdot \Sigma)}_{\text{Overfitting penalty}}.$$

is an unbiased estimator of the MSE of  $\hat{\theta}^\lambda$ .

- Special cases of SURE:
  - Mallows's  $C_p$  (for homoskedastic Gaussian linear regression).
  - Akaike information criterion (for correctly specified unregularized parametric models).
- JS shrinkage can be obtained from Ridge, tuned by minimizing SURE. (Up to a small degrees of freedom correction.)

$$\hat{\theta}^* = \hat{\theta}^{\lambda^*},$$
$$\lambda^* = \underset{\lambda}{\operatorname{argmin}} \operatorname{SURE}(\lambda, \hat{\theta}, \Sigma).$$

## Asymptotic approximations

- Consider large  $n$ , fixed  $k$ ,  $\theta_0$  local to 0.
- Then:
  1. Penalized ERM  $\approx$  shrinkage in the normal means model.
  2. Tuning using  $n$ -fold CV  $\approx$  tuning using SURE.
  3. Out of sample predictive error  $\approx$  MSE.
- **Our main result:**

The risk function of tuned penalized ERM  
converges to  
the MSE of (generalized) JS shrinkage.
- Formally: Under suitable assumptions,

$$E \left[ l(\hat{\theta}_n^{\lambda_n^*} / \sqrt{n}, Z) \right] \rightarrow \frac{1}{2} E[\|\hat{\theta}^* - \theta_0\|^2].$$



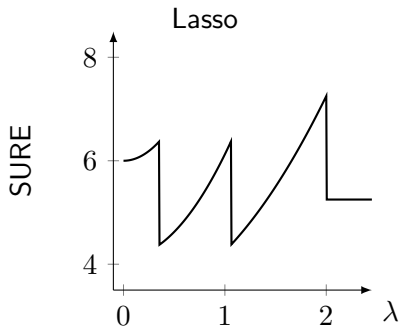
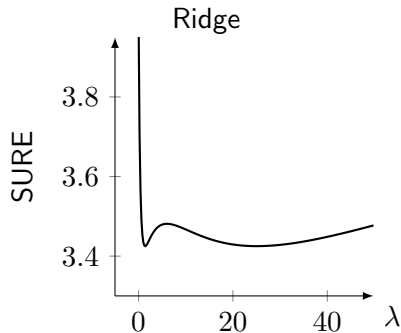
## Key steps

1. Influence function approximations of
  - ERM ( $\Rightarrow$  asymptotically normal),
  - penalized ERM,
  - leave-one-out estimators.
2. Taylor-approximation of CV  
 $\Rightarrow CV \approx SURE$ .
3.  $\Rightarrow$  minimizer of CV  $\approx$  minimizer of SURE
4. Approximation of average out-of-sample loss by squared error.

# Challenges

1. *Convergence of penalized ERM estimators:*  
Standard empirical process results, plus arguments from convex analysis.
2. *Uniformity of convergence of CV to SURE:*  
Need to deal with points of non-differentiability.  
For Lasso: Restrict attention to a grid.
3. *Multimodality of SURE / CV in  $\lambda$ :*  
Non-standard arguments, separately for Ridge and Lasso.  
Show that the problem arises with sufficiently small probability.

## Examples of multi-modality of *SURE*



- Examples for fixed (handpicked) values of  $\hat{\theta}$  and  $\Sigma$ .
- $L^2$  penalty (Ridge) and  $L^1$  penalty (Lasso).

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## Influence function approximations

- Loss:

$$\sum_{i=1}^n l(\theta/\sqrt{n}, Z_n^i) \approx \text{const.} + \frac{1}{2} \|\theta - \tilde{\theta}_n\|^2,$$

where

$$\tilde{\theta}_n = \theta_0 + \frac{1}{\sqrt{n}} \sum_i X_n^i, \quad X_n^i = -\nabla_{\beta} l(\theta_0/\sqrt{n}, Z_n^i).$$

- Asymptotic normality of ERM:

$$\hat{\theta}_n \rightarrow^d \hat{\theta} \sim N(\theta_0, \Sigma).$$

- Penalized ERM:

$$\hat{\theta}_n^{\lambda} \approx \tilde{\theta}_n^{\lambda} = \tilde{\theta}_n + g^{\lambda}(\tilde{\theta}_n).$$

## Influence function approximations for leave-one-out

- Leave-one-out (LOO) loss:

$$\sum_{j \neq i} l(\theta/\sqrt{n}, Z_j^n) \approx \text{const.} + \frac{1}{2} \|\theta - \tilde{\theta}_n^{-i}\|^2,$$

where

$$\tilde{\theta}_n^{-i} = \tilde{\theta}_n - \frac{1}{\sqrt{n}} X_n^i.$$

- Penalized LOO estimator:

$$\begin{aligned} \hat{\theta}_n^{\lambda, -i} &\approx \tilde{\theta}_n^{-i} + g^\lambda(\tilde{\theta}_n^{-i}) \\ &\approx \tilde{\theta}_n^\lambda - \frac{1}{\sqrt{n}} (I + \nabla g^\lambda(\tilde{\theta}_n)) \cdot X_n^i. \end{aligned}$$

Local linear approximation of  $g^\lambda$ : have to be careful for Lasso, which has kinks.

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## Taylor expansion of CV

$$\begin{aligned} CV_n(\lambda) &= \sum_i l(\hat{\theta}_n^{\lambda, -i} / \sqrt{n}, Z_n^i) \\ &\approx \text{const.} + \frac{1}{n} \sum_i \|\hat{\theta}_n^{\lambda, -i} - \theta_0 - \sqrt{n} X_n^i\|^2 \\ &\approx \text{const.} + \frac{1}{n} \sum_i \underbrace{\left\| \tilde{\theta}_n + g^\lambda(\tilde{\theta}_n) - \frac{1}{\sqrt{n}} (I + \nabla g^\lambda(\tilde{\theta}_n)) \cdot X_n^i - \theta_0 - \sqrt{n} X_n^i \right\|^2}_{\approx \hat{\theta}_n^{\lambda, -i}} \\ &\approx \text{const.} + \frac{1}{n} \sum_i \|g^\lambda(\tilde{\theta}_n)\|^2 + \frac{2}{n} \sum_i \langle \nabla g^\lambda(\tilde{\theta}_n) \cdot X_n^i, X_n^i \rangle \\ &\approx \text{const.} + \|g^\lambda(\hat{\theta}_n)\|^2 + 2 \cdot \text{trace}(\nabla g^\lambda(\hat{\theta}_n) \cdot \hat{\Sigma}_n) \\ &= \text{const.} + \text{SURE}(\lambda, \hat{\theta}, \hat{\Sigma}_n), \end{aligned}$$



## Uniform convergence of CV

- Suppose that

$$\lim_{\|\delta\| \rightarrow 0} \sup_{\lambda \in \Lambda} \frac{\|R^\lambda(\delta; \theta)\|}{\|\delta\|} = 0$$

for (Lebesgue) almost every  $\theta$ , where

$$R^\lambda(\delta; \theta) = g^\lambda(\theta + \delta) - g^\lambda(\theta) - \nabla g^\lambda(\theta) \cdot \delta.$$

- Then (assuming regularity conditions) the n-fold crossvalidation criterion satisfies

$$\sup_{\lambda \in \Lambda} \left| CV_n(\lambda) - SURE(\lambda, \hat{\theta}_n, \Sigma) \right| \xrightarrow{P} 0.$$

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## Dealing with multi-modality: Ridge

- Penalty  $\pi(\theta) = \frac{1}{2}\theta \cdot A^{-1} \cdot \theta$ , where  $A$  is positive definite.  $\Rightarrow$

$$g^\lambda(\theta) = C_\lambda \cdot \theta, \quad \nabla g^\lambda(\theta) = C_\lambda, \quad C_\lambda = \left(\frac{1}{\lambda}A + I\right)^{-1}.$$

- Thus

$$SURE(\lambda, \theta, \Sigma) = \text{trace}(\Sigma) + \|C_\lambda \cdot \theta\|^2 + 2 \text{trace}(C_\lambda \cdot \Sigma).$$

- Change of coordinates, with slight abuse of notation:  
For  $R = \|\hat{\theta}\|$  and  $\nu = \hat{\theta}/R$ ,

$$SURE(\lambda, R, \nu) := SURE(\lambda, \hat{\theta}, \Sigma).$$

## SURE for Ridge

1. For every  $\theta$ ,

$$\lim_{\theta' \rightarrow \theta} \sup_{\lambda} |SURE(\lambda, \theta', \Sigma) - SURE(\lambda, \theta, \Sigma)| = 0$$

2.  $SURE(\lambda, R, \nu)$  is strictly **supermodular** in  $\lambda$  and  $R$ . This implies:
  - 2.1  $\lambda(R, \nu) = \operatorname{argmin}_{\lambda \in \mathbb{R}^+} SURE(\lambda, R, \nu)$  is **monotonically decreasing** in  $R$ , given  $\nu$ .
  - 2.2  $\lambda(R, \nu)$  has at most **countably many discontinuities**, as a function of  $R$ , given  $\nu$ .
3. Fix  $\nu$  and  $R$  such that  $\lambda(\cdot)$  is continuous in  $R$  at  $(R, \nu)$ , and let  $\bar{\lambda} = \lambda(R, \nu)$ . Then supermodularity implies that **the minimum of SURE is well separated**: For any  $\epsilon > 0$ ,

$$\inf_{\lambda \in \mathbb{R}^+ \setminus [\bar{\lambda} - \epsilon, \bar{\lambda} + \epsilon]} SURE(\lambda, R, \nu) - SURE(\bar{\lambda}, R, \nu) > 0.$$

## Almost everywhere continuity for Ridge

- Define

$$w = (\theta, \Delta), \quad \|w\| = \|\theta\| + \sup_{\lambda} |\Delta(\lambda)|$$

- For **almost every**  $\theta$ , the mapping from  $w$  to

$$g^{\tilde{\lambda}(\theta, \Delta)}(\theta) = \left( \frac{1}{\tilde{\lambda}(\theta, \Delta)} A + I \right)^{-1} \cdot \theta, \text{ where}$$
$$\tilde{\lambda}(\theta, \Delta) = \min_{\lambda} \left( \operatorname{argmin}_{\lambda} [SURE(\lambda, \theta, \Sigma) + \Delta(\lambda)] \right),$$

is **continuous at**  $w = (\theta, 0)$  with respect to the norm  $\|w\|$ .

- Continuous mapping theorem, uniform integrability  
 $\Rightarrow$  convergence of risk.

## Dealing with multi-modality: Lasso

- Penalty  $\pi(\theta) = \|A^{-1} \cdot \theta\|_1$ , where  $A$  is an invertible matrix.
- Denote  $h^\lambda(\theta) = A^{-1}(\theta + g^\lambda(\theta)) \Rightarrow$

$$h^\lambda(\theta) = \underset{h}{\operatorname{argmin}} \frac{1}{2} \|A \cdot h - \theta\|^2 + \lambda \cdot \|h\|_1.$$

- Solution:

$$h_J^\lambda(\theta) = (A_J' A_J)^{-1} \cdot [A_J' \theta - \lambda \eta_J],$$

where

- $\eta_j = \operatorname{sign}(h_j^\lambda)$ ,
- $J = \{j : \eta_j \neq 0\}$ .

## SURE for Lasso

1. As a function of  $\lambda$ , the graph of  $SURE(\lambda, R, \nu)$  consists of **at most  $3^k$  segments** on which  $\eta$  and  $J = \{j : \eta_j \neq 0\}$  are constant, and

$$SURE(\lambda, R, \nu) = \text{const.} + \lambda^2 \cdot \eta'_J (A'_J A_J)^{-1} \eta_J.$$

2. Let  $\lambda_1, \lambda_2, \dots, \lambda_m$  ( $m \leq 3^k$ ) be the local minimizers of  $SURE(\lambda, 1, \nu)$ . Then  $R \cdot \lambda_1, R \cdot \lambda_2, \dots, R \cdot \lambda_m$  are the **local minimizers** of  $SURE(\lambda, R, \nu)$ .
3. Let  $\lambda(R, \nu) = \operatorname{argmin}_{\lambda \in \mathbb{R}^+} SURE(\lambda, R, \nu)$ . Then
  - $\lambda(R, \nu) = R \cdot \lambda_{j(R)}$ ,
  - where  $j(R) \in \{1, 2, \dots, m\}$  is a **monotonically decreasing**.
4. Fix  $\nu$  and  $R$  such that  $\lambda(\cdot)$  is continuous in  $R$  at  $(R, \nu)$ , and let  $\bar{\lambda} = \lambda(R, \nu)$  be such that  $\eta \neq 0$ . Then **the minimum of SURE is well separated**: For any  $\epsilon > 0$ ,

$$\inf_{\lambda \in \mathbb{R}^+ \setminus [\bar{\lambda} - \epsilon, \bar{\lambda} + \epsilon]} SURE(\lambda, R, \nu) - SURE(\bar{\lambda}, R, \nu) > 0.$$

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# Conclusion

- Majority of supervised learning methods:
  - Empirical risk minimizers
  - with regularization
  - tuned using cross-validation.
- We show:
  - Such methods behave approximately like (generalized) James-Stein shrinkage:
  - Uniform dominance relative to un-regularized estimators.
  - Largest gains for: Large  $k$ , small  $\|\theta\|$ .

## Open issues and limitations

- Our asymptotic result for Lasso holds for a fixed grid  $\Lambda$ .
  - Extension to sequence of grids?
  - More refined argument to cover the case  $\Lambda = \mathbb{R}$ ?
- Our approximations hold for fixed  $k$ , large  $n$ .
  - $\Rightarrow$  We can leverage asymptotic normality.
  - But what about the over-parametrized case  $k > n$ ?  
Important in deep learning!
- Risk  $\approx$  average loss for point prediction.
  - Conformal inference:  
Turns point prediction into predictive intervals with guaranteed coverage.
  - Can we map our risk results into results about average size of predictive sets?

Thank you!