

A model of multiple hypothesis testing

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 - Separate testing generally does not control notions of **compound error** at 5%.
- There is substantial variation on the choice of compound error and/or tests
 - Family-wise error rate (**FWER**): probability of rejecting at least one true null;
 - False discovery rate (**FDR**): expected proportion of incorrectly rejected null hypotheses;
 - **Indexing**: aggregate outcomes into a single index [e.g., Anderson (2008)].
- Several algorithms to control compound errors (e.g., Bonferroni correction)

Multiple subgroups in clinical trials

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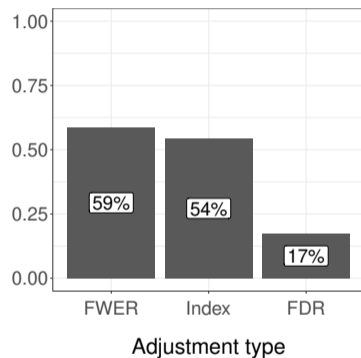
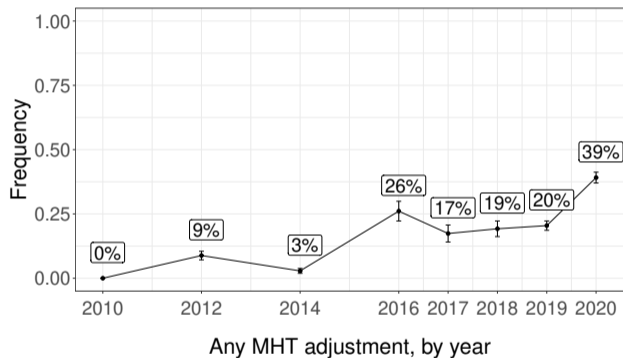
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“This observed heterogeneity led two regulatory agencies to different assessments. The National Institute for Health and Care Excellence (NICE, English and Welsh authority) concluded a clinical benefit for the overall population whereas the Institut für Qualität und Wirtschaftlichkeit im Gesundheitswesen (IQWiG, German authority) concluded efficacy only for the most beneficial subgroup of patients (symptomatic peripheral arterial disease)” (Tanniou et al., 2016)

Policy experiments with multiple treatments in Economics: top-5 journals



Heterogeneity in when and how to adjust inference.

This paper

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- Model's assumptions can justify single-hypothesis testing [Tetenov (2016)]

We study MHT for different types of multiplicity

Key feature of our model: hypothesis tests correspond to policy decisions

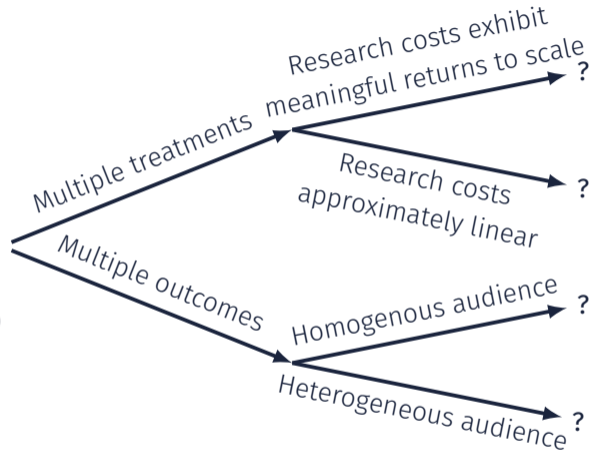
- **Multiple treatments (or subgroups):** simple mapping betw/ tests and decisions
- **Multiple outcomes:** might or might not interpret as informing multiple decisions
 - Research informs a **single policy decision** (e.g., whether to scale up an intervention)
 - Research informs **multiple heterogeneous policy-makers**

Outline

1. Multiple treatments

2. Multiple outcomes (one treatment)

3. Empirical Analysis and conclusions



- Economic analysis of optimal statistical approaches [E.g., Chassang et al. (2012); Tetenov (2016); Spiess (2018); Henry and Ottaviani (2019); Di Tillio et al. (2017); Kasy and Spiess (2023)]
 - We focus on MHT
- Models of scientific communication [E.g., Frankel and Kasy (2022); Andrews and Shapiro (2021); Banerjee et al. (2017)]
 - We relate the structure of the scientific process to MHT
- Work on decision theory and hypothesis testing [E.g., Wald (1950); Robbins (1951); Storey (2003); Lehmann and Romano (2005); Efron (2008)]
 - We provide an economic model with incentives that allows for discriminating between different MHT procedures. We show when MHT is optimal and when it is not.
- Statistical methods for MHT corrections [E.g., Holm (1979); Westfall and Young (1993); Benjamini and Hochberg (1995); Romano et al. (2010)]
 - We provide guidance for choosing appropriate methods

Multiple treatments (interventions or subgroups)

Players and stakeholders' welfare

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- ⇒ **Policy implementation**
- upon experimentation, **additive welfare effects** $\theta^\top r(X)$ on stakeholders (no spillovers)
 - Later: settings with interactions between treatments

Example

Ex **Parameters of interest:** the researcher evaluates J treatments D_1, \dots, D_J using

$$Y = \theta_1 D_1 + \dots + \theta_J D_J + \varepsilon$$

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Ex **Testing protocol (t -test):** $r_j(X) = 1\{X_j / \sqrt{\Sigma_{j,j}} > t\}$ for $j = 1, \dots, J$

\Rightarrow In the paper general testing protocol

Game

Stage 1: the social planner, who **doesn't know** θ , chooses r to maximize worst-case welfare: for $\lambda \geq 0, \pi \in \Pi$

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Stage 2: given r , the researcher, **who knows** θ (can be relaxed), experiments if her expected utility $\beta_r(\theta)$ is positive, where

$$\beta_r(\theta) = \underbrace{\int \sum_{j=1}^J r_j(x) dF_\theta(x)}_{\substack{\text{benefit from approval} \\ \text{(expected number of rejections)}}} - \underbrace{C(J)}_{\text{research costs relative to benefits}}$$

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 - ⇒ Pre-specification typically recommended
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Some extensions (main conclusions unchanged):

- Sub-populations have varying size
- The researcher has a prior over θ
- Endogenous choice of which treatment to test (and J), but pre-specify

Characterization of maximin protocols ($\lambda = 0$)

- **Proposition:** r^* is maximin optimal if and only if

$$(a) \beta_{r^*}(\theta) \leq 0 \quad \forall \theta \in \Theta_0 \quad \text{and} \quad (b) \quad v_{r^*}(\theta) \geq 0 \quad \forall \theta \in \Theta \setminus \Theta_0$$

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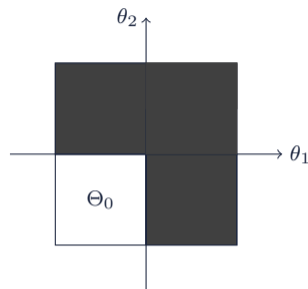
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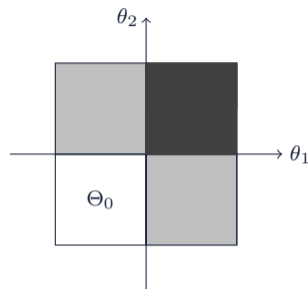
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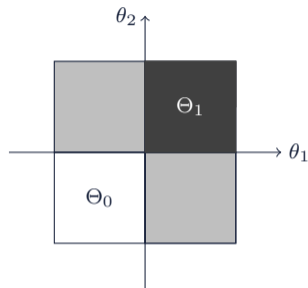
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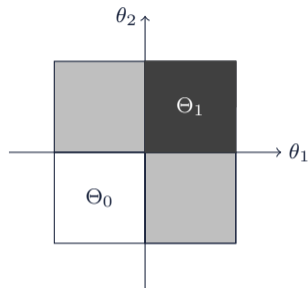
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- **Intuition**

- When $\theta \in \Theta_0$, research has only downside \Rightarrow keep approval probability low
- When the cost doesn't depend on J , this condition will be violated for large enough J

Optimal protocols ($\lambda \geq 0$)

- There are many maximin protocols (including $r_j(X) = 0$ for all j).
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- (i) r^* is maximin optimal (\implies size control/non-negative welfare)
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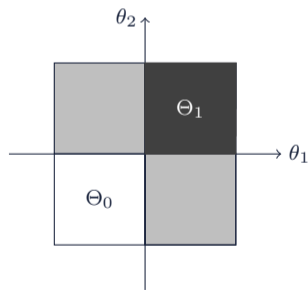
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⇒ Challenge here to show maximin optimality in mixed orthants

Optimal MHT adjustments depend on the research costs

- Decompose the costs into **fixed costs** and **variable costs**: $C(J) = c_f + c_v(J)$
- Optimal level for separate t -tests: $\alpha(J) = (c_f + c_v(J))/J$
- **Examples**

	Cost function	Level	Intuition
Bonferroni	$c_f = \alpha, c_v(J) = 0$	α/J	Adjustment for increased benefits due to false positives
No adjustment	$c_f = 0, c_v(J) = \alpha J$	α	MHT adjustments are “built into” the cost structure

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	Cost function	Level	Intuition
Bonferroni	$c_f = \alpha, c_v(J) = 0$	α/J	Adjustment for increased benefits due to false positives
No adjustment	$c_f = 0, c_v(J) = \alpha J$	α	MHT adjustments are “built into” the cost structure

- General MHT adjustment based on relative costs:

$$\alpha(J) = \underbrace{\frac{C(J)/J}{C(1)}}_{\text{adjustment factor}} \times \alpha(1)$$

Additional formal results in the paper

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 - ⇒ maximin optimality using worst-case upper bounds
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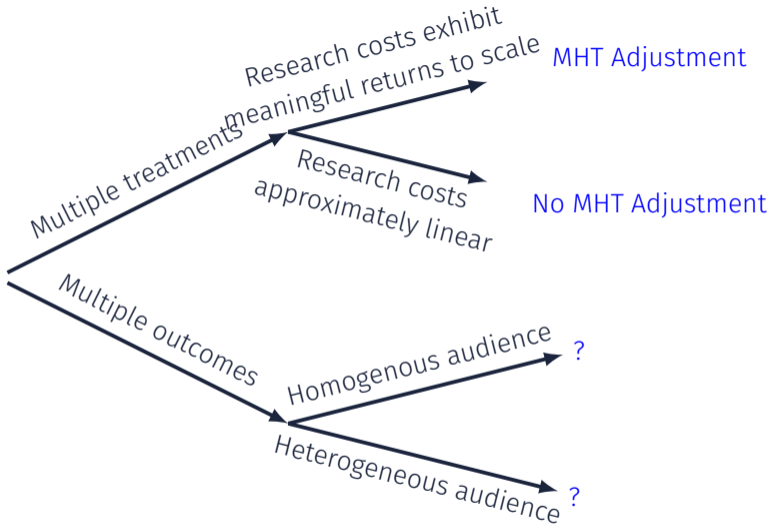
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- **Other notions of power**
 - Worst-case power: study a ϵ deviations from the positive orthant
 - Weighted Average Power: no rule most powerful for any choice of the weights



Multiple outcomes (one treatment)

- There are G outcomes $Y = (Y_1, \dots, Y_G)$ associated with $X = (X_1, \dots, X_G)$
- **Example:** for $g = 1, \dots, G$, the researcher estimates the effect of treatment D on outcome Y_g using the regression model $Y_g = \mu + \theta_g D + \varepsilon_g \Rightarrow X = (\hat{\theta}_1, \dots, \hat{\theta}_G)^\top$

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- There is an audience of J policy-makers each with individual utility $u_j(\theta)$
- Researcher makes J recommendations, one for each policy-maker:

$$r(X) = (r_1(X), \dots, r_J(X))$$

Multiple policymakers ($G = J$)

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- Isomorphic to model with multiple treatments, and optimal t -tests

$$r_j^*(X) = 1 \left\{ \frac{X_j}{\sqrt{\Sigma_{j,j}}} \geq \Phi^{-1} \left(1 - \frac{C(G)}{G} \right) \right\}, \quad \forall j.$$

Single policymaker

- Here $r(X) \in \{0, 1\}$, $\beta_r(\theta) = \int r(x)dF_\theta(x) - C(G)$, $u(\theta) = \theta^\top w^*$

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$$r^*(X) = 1 \left\{ \frac{X^\top w^{\min}}{\sqrt{w^{\min\top} \Sigma w^{\min}}} > \Phi^{-1}(1 - C(G)) \right\},$$

where w^{\min} minimizes $\sqrt{w^\top \Sigma w}$ st $\sum_g w_g = 1$ (**Statistical aggregation**)

Empirical studies

Clinical trials

- Sertkaya et al. (2016) estimate that 46% costs are fixed in average Phase 3 trial
- Take cost function $C(J) = c_f + mJ$ satisfying $c_f/(c_f + m\bar{J}) = 0.46$, where \bar{J} is the number of subgroups in a typical study
- Take $\bar{J} = 3$ based on Pocock et al. (2002) implying $\alpha(J) = \alpha(1) \times \left[\frac{1+2.56/J}{3.56} \right]$

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J	$\alpha(1) = 0.025$	$\alpha(1) = 0.05$	$\alpha(1) = 0.1$	$\alpha(1) = 0.15$
1	0.025	0.050	0.100	0.150
2	0.016	0.032	0.064	0.096
3	0.013	0.026	0.052	0.078
4	0.012	0.023	0.046	0.069
5	0.011	0.021	0.042	0.064
9	0.009	0.018	0.036	0.054
∞	0.007	0.014	0.028	0.042

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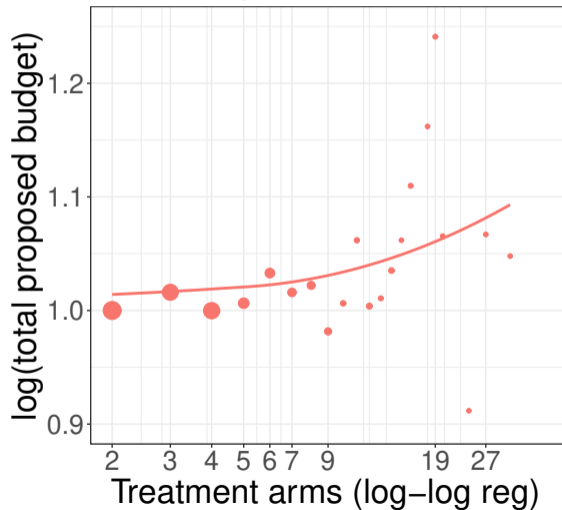
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Summary of the results

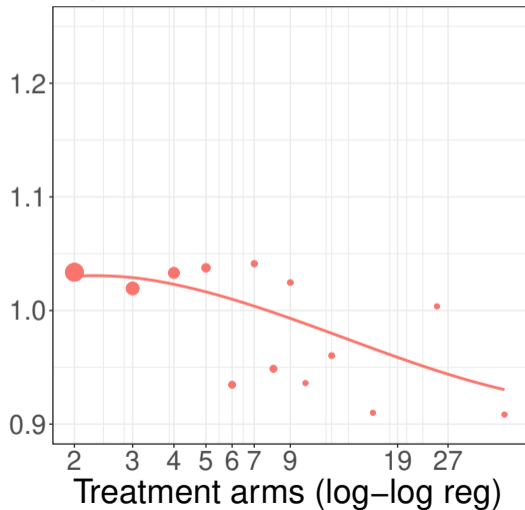
- Returns to scale with number of arms \Rightarrow some MHT adjustments are needed
- Costs are *not* invariant to scale \Rightarrow Bonferroni is too stringent
- Costs vary with context \Rightarrow in high-income countries, studies with more treatment arms are also the cheaper (may reflect different research technology)

Data visualization

Main Sample



High income countries



Results

	Main sample		
	(1)	(2)	(3)
log(Treatment Arms) [β]	0.180 (0.077)	0.183 (0.064)	0.215 (0.080)
Proposal Type FEs	No	Yes	Yes
Initiative FEs	No	No	Yes
p -value, $H_0 : \beta = 0$	0.019	0.004	0.007
p -value, $H_0 : \beta = 1$	0.000	0.000	0.000
Observations	812	812	655
Adjusted R ²	0.005	0.352	0.380

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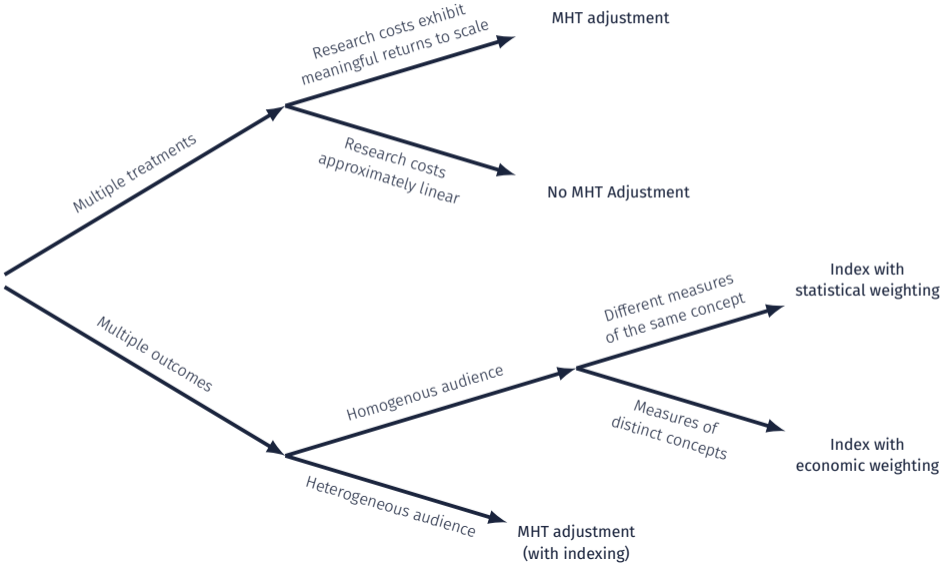
- Taking $\hat{\beta} \approx 0.2$ for the main sample implies $\alpha(J) = \alpha(1)J^{0.2-1}$

Conclusions

Extensions

- Endogenous J (pre-specified by the researcher ex-ante)
- Unknown θ and researcher's prior on θ
- Some benevolent researcher
- Additional forms of interactions
- Alternative notions of power (WAP and local power)
- Variance that might depend on J and heterogeneous variance
- Weighted welfare function
- Two sided tests

Conclusions



Thank you!

Questions? Thoughts? Comments? Please reach out:
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`pniehaus@ucsd.edu`

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Digression: Can we justify the FDR?

- Suppose that $u_j(\theta) = \theta_j$. FDR is optimal if

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- FDR does not arise as a natural solution in our frequentist maximin framework
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